Finite volume and asymptotic methods for stochastic neuron models with correlated inputs

Robert Rosenbaum, Jianfu Ma, Fabien Marpeau, Aditya Barua and Krešimir Josić

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Abstract

We consider a pair of stochastic integrate and fire neurons receiving correlated stochastic inputs. The evolution of this system can be described by the corresponding Fokker-Planck equation with non-trivial boundary conditions resulting from the refractory period and firing threshold. We propose a finite volume method that is orders of magnitude faster than the Monte Carlo methods traditionally used to model such systems. The resulting numerical approximations are proved to be accurate, nonnegative and integrate to 1. We also approximate the transient evolution of the system using an Ornstein–Uhlenbeck process, and use the result to examine the properties of the joint output of cell pairs. The results suggests that the joint output of a cell pair is most sensitive to changes in input variance, and less sensitive to changes in input mean and correlation.

1 Introduction

The integrate and fire (IF) model is used widely in mathematical biology (Burkitt, 2006; Keener and Sneyd, 2008). It is simple, yet versatile, and provides a good approximation of the response of an excitable cell in a variety of situations. A stochastic version of the IF model can describe the behavior of large populations of cells through the evolution of the corresponding probability density (Knight, 1972; Nykamp and Tranchina, 2000; Rolls et al., 2008). It can also be used to study the response of a single cell subject to a large number of small, statistically independent inputs (Lindner, 2001; Renart et al., 2003).

Collections of excitable cells frequently do not behave independently. The joint response of populations of electrically active cells is of interest in a number of areas in biology: Pancreatic β-cells have to synchronize their response to secrete insulin (Meda et al., 1984; Sherman and Rinzel, 1991), and the coordinated activity of cardiac cells is essential for their function (Keener and Sneyd, 2008). Our study is motivated primarily by cells in neural populations. Such cells typically fire action potentials in response to synaptic inputs from other cells. The standard stochastic IF model can capture the response of a cell when such inputs are independent (Renart et al., 2003). However dependencies between these inputs cannot always be ruled out. Such dependencies can affect the output statistics of a neuronal population, and significantly impact the amount of information carried in the population response (Salinas and Sejnowski, 2000; Sompolinsky et al., 2001). Even weak correlations between individual cells can significantly impact the ensemble activity of a population (Shadlen and Newsome, 1998; Rosenbaum et al., 2010). Here we describe a model that can be used to examine the behavior of two cell populations (or cell pairs) receiving correlated inputs.

We first develop a Fokker-Planck equation that describes the evolution of the probability density for a pair of cells receiving correlated inputs. The response of cell pairs receiving correlated inputs has been studied previously using linear response theory (Ostojić et al., 2009; de la Rocha et al., 2007), and numerical simulations (Galán et al, 2007) in related models. However, the boundary conditions in the presence of a refractory period are nontrivial, and can impact the behavior of the system. We therefore present the model in some detail.

Next we describe a finite volume method that can be used to study solve the Fokker-Planck equation numerically for the probability density. Previously, we proposed a fast and accurate finite volume method for modeling a general IF neuron driven by a stochastic input (Marpeau et al., 2009). This method was significantly faster than Monte Carlo (MC) simulations and we proved several stability properties of the algorithm. Here we extend this method to two neurons

with correlated inputs. While the dynamics of interacting populations has been examined previously (Nykamp and Tranchina, 2000; Harrison, et al 2005), we are not aware of a numerical treatment of the Fokker–Planck equation corresponding to stochastic IF neurons driven by correlated noise.

Finally, we develop a simple analytical approximation in terms of a related Ornstein-Uhlenbeck process that captures the response of a cell pair to study the behavior of cells, or cell populations receiving correlated inputs. This approach provides an alternative to the linear response techniques commonly in use (Lindner, 2001; Ostojić et al., 2009).

We use this approximation to examine the response of a single cell and a cell pair to changes in the input parameters. The variance (noisiness), mean and synchrony between the inputs are separate channels along which information can be communicated to postsynaptic cells. We find that the spiking statistics of a single cell and the cell pair are most sensitive to changes in the variance of the input. This suggests that the joint response of a cell population most accurately tracks input noise intensity.

2 Model Description

A single IF neuron with stochastic input is described by the Langevin equation:

$$\frac{dV}{dt} = f(V) + \sqrt{2D}\xi(t), \qquad V \in (-\infty, V^T). \tag{1}$$

Here f defines the deterministic (drift) behavior, and $\xi(t)$ a Gaussian stochastic processes with $\langle\xi(t)\rangle = 0$ and $\langle\xi(t)\xi(t')\rangle = \delta(t - t')$. When the voltage reaches a threshold, V^T, a spike is fired, and V is instantaneously reset to $V^R < V^T$. A spike may be followed by an absolute refractory period τ, during which a neuron is insensitive to inputs, and V is held fixed at V^R.

This model can also be understood as a the diffusive limit of a population of cells receiving independent inputs (Omurtag, 2000). To model a pair of cells receiving correlated inputs, we assume that their membrane voltages V and W obey a pair of Langevin equations:

$$\begin{aligned}
\dot{V} &= f(V, W) + I_V(t); \quad I_V(t) = \mu_V + \sqrt{2D}(\sqrt{1-c}\xi_V(t) + \sqrt{c}\xi_c(t)) \\
\dot{W} &= g(W, V) + I_W(t); \quad I_W(t) = \mu_W + \sqrt{2D}(\sqrt{1-c}\xi_W(t) + \sqrt{c}\xi_c(t)).
\end{aligned} \tag{2}$$

The inputs, $I_j(t)$, received by the cells are comprised of statistically *independent* stochastic processes $\xi_V(t)$ and $\xi_W(t)$, and a common input $\xi_c(t)$. The ξ_i are again assumed to be Gaussian with $\langle\xi_i(t)\rangle = 0$ and $\langle\xi_i(t)\xi_j(t')\rangle = \delta(t - t')\delta_{i,j}$. The constant c, is the Pearson correlation coefficient between the inputs and lies between 0 and 1. For instance, for two leaky integrate-and-fire (LIF) neurons with common input, but no direct coupling, $f(V, W) = -g_V(V - V_{\text{rest}})$ and $g(V, W) = -g_W(W - W_{\text{rest}})$. Each cell spikes when the voltage V crosses the threshold, V^T and W^T respectively. After each spike the voltage is reset to $V^R < V^T$ ($W^R < W^T$ for cell 2), and is pinned to this value for the duration of the refractory period, τ_V (τ_W for the second cell, see Fig. 1). For simplicity we will refer to the two as neuron V and W, although this can be understood as "populations V and W" (Harrison, et al 2005). The joint probability density of the two voltages evolves on the domain $\Omega = (V^{-\infty}, V^T) \times (W^{-\infty}, W^T)$. In theoretical studies it is frequently assumed that $V^{-\infty} = W^{-\infty} = -\infty$. However, since we will be interested in numerical simulations, we assume that these quantities are large and negative.

With $U = (V, W)$ and $F = (f, g)$ the Fokker-Planck equation corresponding to Eq. (2) takes the form

$$\partial_t P(t, U) + \text{div}\left(F(U)P(t, U) - DM\nabla P(t, U)\right) = 0, \qquad M = \begin{pmatrix} 1 & c \\ c & 1 \end{pmatrix}, \tag{3}$$

for $V \in (V^{-\infty}, V^T) \backslash V^R$, $W \in (W^{-\infty}, W^T) \backslash W^R$. Here D is the diffusion coefficient and M is the correlation matrix. This equation is coupled with reflecting boundary conditions at $V = V^{-\infty}$ or $W = W^{-\infty}$, and absorbing boundary conditions at both thresholds:

$$\begin{aligned}
\left(f(U) - D\partial_V - cD\partial_W\right)P(t, U)\big|_{V=V^{-\infty}} &= 0, \qquad \left(g(U) - D\partial_W - cD\partial_V\right)P(t, U)\big|_{W=W^{-\infty}} = 0, \\
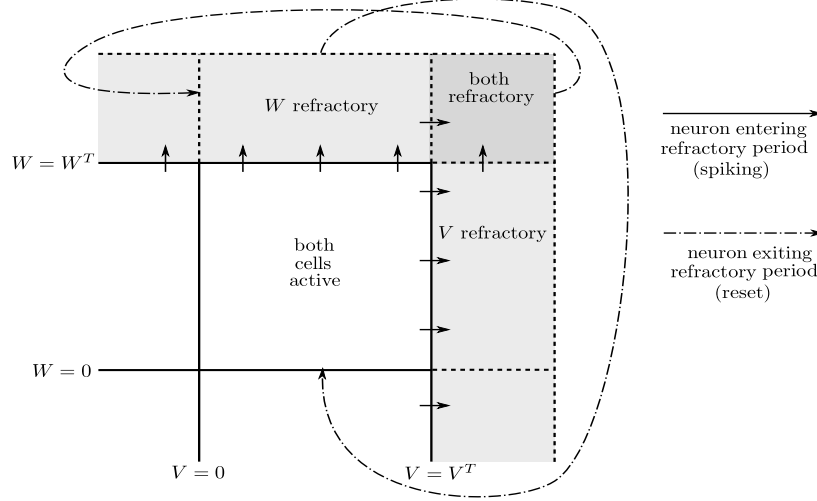P(t, U)\big|_{V=V^T} &= P(t, U)\big|_{W=W^T} = 0.
\end{aligned} \tag{4}$$

Figure 1: (Left) Domain of simulation. (Right) Circulation of probability mass through populations P, R_V, R_W and R.

The presence of the refractory behavior in the IF model introduces additional complexity. If either neuron enters its refractory state, the corresponding voltage is fixed at the reset value, and the entire system effectively evolves according to a one-dimensional Fokker-Planck equation. During this time, it is possible that the second cell also crosses the threshold, fires and enters the refractory state. In this case the voltages are fixed at (V^R, W^R) and both neurons are insensitive to inputs until one of them exits the refractory state.

To capture the behavior of neurons in the refractory period, we model the evolution of densities using three separate, communicating sub-populations, in addition to $P(t, U)$ (Sirovich, 2008; Ly and Tranchina, 2009):

- $R_V(t, r, V)$, the probability density of the fraction of the population in which only neuron W is in the refractory state,

- $R_W(t, s, W)$ the corresponding density in which only neuron V is in the refractory state, and

- $R(t, s, r)$, the density corresponding to both neurons in the refractory state.

For all densities, t refers to the time since the beginning of the simulation, while r and s refer to relative times measured from the beginning of the refractory period for neuron V and W, respectively. Therefore, $R_V(t, r, V_0)\Delta V \Delta r$ is the fraction of the population for which neuron W has been in the refractory period between r and $r + \Delta r$ units of time, and the voltage of neuron V is between V_0 and $V_0 + \Delta V$. The quantity $R(t, s, r)\Delta s \Delta r$ is the fraction of the population in which neurons V and W have been in refractory periods for times in the intervals $[s, s + \Delta s]$ and $[r, r + \Delta r]$ respectively. The use of variables s and r is closely related to age-structured population dynamics models (Iannelli, 1994; Webb, 1985). Indeed, s and r denote the "ages" of the refractory states for neurons 1 and 2 respectively. Fig. 1 summarizes the circulation of probability mass between the different populations involved.

Since the entire population is described by these densities we have for any time t

$$\int_{V-\infty}^{V^T} \int_{W-\infty}^{W^T} P(t, V, W)\, dW\, dV + \int_0^{\tau_W} \int_{W-\infty}^{W^T} R_W(t, s, W)\, dW\, ds$$
$$+ \int_0^{\tau_V} \int_{V-\infty}^{V^T} R_V(t, s, V)\, dV\, ds + \int_0^{\tau_V} \int_0^{\tau_W} R(t, s, r)\, ds\, dr = 1. \quad (5)$$

We next describe the evolution of the main population and the three refractory populations and how they are coupled to each other through boundary terms. The refractory populations evolve according to one-dimensional Fokker-

Planck equations,

$$(\partial_t + \partial_r)R_V(t,r,V) + \mathrm{div}\left(f(U)R_V(t,r,V) - D\partial_V R_V(t,r,V)\right) = 0, r \in (0,\tau_W),$$
$$(\partial_t + \partial_s)R_W(t,s,W) + \mathrm{div}\left(g(U)R_W(t,s,W) - D\partial_W R_W(t,s,W)\right) = 0, s \in (0,\tau_V), \tag{6}$$

$$(\partial_t + \partial_s + \partial_r)R(t,s,r) = 0. \tag{7}$$

In addition, mass from the main population $P(t,U)$ is injected into populations $R_V(t,r,V)$ and $R_W(t,s,W)$ at $r=0$ and $s=0$ as neuron V and W cross threshold respectively. These source terms are described by

$$R_V(t,0,V) = -D\partial_W P(t,V,W^T), \; R_W(t,0,W) = -D\partial_V P(t,V^T,W) . \tag{8}$$

While either neuron is in the refractory state, the other neuron can enter its own refractory state as well, providing boundary conditions to equation (7) for inward characteristics:

$$R(t,s,0) = -D\partial_W R_W(t,s,W^T), \; R(t,0,r) = -D\partial_V R_V(t,r,V^T) . \tag{9}$$

Next, while both neurons are in the refractory state, neuron V or W may exit the refractory state while the other neuron remains in the refractory state. Using $[\xi]\big|_{z=Z} := \lim_{z \to Z^+} \xi(z) - \lim_{z \to Z^-} \xi(z)$ to denote the jump across point $Z \in \mathbb{R}$, we can express the contribution of population $R(t,s,r)$ to the following source terms (Melnikov, 1993; Lindner, 2001)

$$\left[D\partial_V R_V(t,s,.)\right]\Big|_{V=V^R} = R(t,\tau_V,s), \qquad \left[D\partial_W R_W(t,r,.)\right]\Big|_{W=W^R} = R(t,r,\tau_W). \tag{10}$$

As either neuron exits the refractory period, it re-enters the main population modeled by the density $P(t,U)$. This is captured by adding the source terms:

$$\left[D\partial_V P(t,.,V)\right]\Big|_{V=V^R} = R_W(t,\tau_V,W), \qquad \left[D\partial_W P(t,V,.)\right]\Big|_{W=W^R} = R_V(t,\tau_W,V) . \tag{11}$$

The densities are also continuous across the reset potentials, so that

$$\left[P(t,.,W)\right]\Big|_{V=V^R} = \left[P(t,W,.)\right]\Big|_{W=W^R} = \left[R_V(t,s,.)\right]\Big|_{V=V^R} = \left[R_W(t,s,.)\right]\Big|_{W=W^R} = 0 . \tag{12}$$

Finally, reflecting and absorbing boundary conditions are imposed on Eq. (6), and (7) by requiring:

$$\left(f(U) - D\partial_V\right)R_V(t,s,.)_{|V=V^{-\infty}} = \left(g(U) - D\partial_W\right)R_W(t,s,.)_{|W=W^{-\infty}} = 0 , \tag{13}$$

$$R_V(t,s,V^T) = R_W(t,s,W^T) = 0 . \tag{14}$$

Notice that when there is no refractory period ($\tau_V = \tau_W = 0$), (8)–(11) reduce to the single boundary condition

$$\left[D\partial_V P(t,.,W)\right]_{|V=V^R} = -D\partial_V P(t,V^T,V) , \; \left[D\partial_W P(t,V,.)\right]_{|W=W^R} = -D\partial_W P(t,V,W^T) .$$

3 Description of the Numerical Methods

The numerical methods used in simulating the solutions of the model described in the previous section are not completely standard. The anisotropy of the diffusion operator coupled with the absorbing boundary condition presents numerical challenges different from those encountered, for instance, when modeling phase oscillators (Galán et al, 2007). We therefore give a brief description of our approach here.

4

3.1 Finite Volume Method

Three requirements in the numerical discretization of Eqs. (3)–(14) are obtaining numerical probability densities which are accurate, nonnegative and integrate to 1. We dealt with similar difficulties in (Marpeau et al., 2009), and use an extension of that approach here. To obtain numerical densities which integrate to 1, we use conservative numerical schemes which ensure that the mass lost by a mesh-element is transmitted exactly to its neighboring elements. This ensures preservation of mass of the initial densities. Secondly, the drift operator is well known to be an obstacle when trying to combine accuracy, non-negativity, and stability of the numerical densities. Therefore, we use an operator splitting method (described below) that enables us to discretize the drift and diffusion operators separately. The discretization of the drift term is carried out by an upwind scheme whose accuracy is improved using flux limiters.

However, the two-dimensional nature of the problem induces further difficulties:

1. Extra numerical diffusion is generated in the cross directions from the drift operator, leading to a loss of accuracy.

2. Due to the correlation coefficient c, the diffusion matrix DM is anisotropic. The discretization of the cross derivatives $\partial_{VW}^2 P$ commonly involves the inversion of matrices that are not unconditionally strongly diagonally dominant, which makes it difficult to obtain nonnegative numerical densities.

3. Due to the refractory periods, the multiple boundary conditions (6)–(14) drastically increase the algorithmic complexity, compared with the one-dimensional model in (Marpeau et al., 2009). In particular, the densities R_V and R_W have to be obtained by discretizing Eq. (6) at each time step and age step. Moreover, the code has to gather all the phenomena and assemble the circulation between the four populations in an efficient manner.

The technical of our approach are relegated to Appendix A, and what follows is an outline. First, the time interval $\mathbb{R}_+$ on which the solution will be approximated is partitioned into sub-intervals (t_n, t_{n+1}). We will denote the numerically obtained approximation of the population P at time t_n by P^n. As in (Marpeau et al., 2009), P^{n+1} is obtained from P^n by splitting equations (3)–(4) into

$$\partial_t P + \mathrm{div}\,(fP) = 0, \; P\big|_{V=V^{-\infty}} = P\big|_{W=W^{-\infty}} = 0, \tag{15}$$

$$\partial_t P - \mathrm{div}\,(DM\nabla P) = 0,$$
$$\left(D\partial_V + CD\partial_{VW}\right)P\big|_{V=V^{-\infty}} = \left(D\partial_W + CD\partial_{VW}\right)P\big|_{W=W^{-\infty}} = P\big|_{V=V^{-\infty}} = P\big|_{W=W^{-\infty}} = 0. \tag{16}$$

The numerical solution is updated at time t_{n+1} using

$$P^{n+1} = \mathscr{S}_2\big(\mathscr{S}_1(P^n)\big), \tag{17}$$

where $\mathscr{S}_1$ and $\mathscr{S}_2$ are approximation schemes for Eq. (15) and (16) respectively, along with split interior conditions specified later. This technique allows us to develop specific numerical schemes which are adapted to each differential operator in Eq. (3).

The numerical scheme $\mathscr{S}_1$ is nonlinear explicit. A compromise between accuracy and stability is obtained by adding flux limiters to the upwind scheme. As discussed in the Appendix, we use the numerical scheme introduced in (Marpeau et al., 2009) in each direction, V and W. The time step is restricted by a Courant-Friedrichs-Lewy (CFL) condition (Courant, 1928; Godlewski and Raviart, 1990), which provides stability and positivity preservation of the scheme by ensuring that the drift term does not shift the numerical solution by more than one mesh element per time step (Godlewski and Raviart, 1990).

The scheme $\mathscr{S}_2$ is linear implicit. Using centered approximation of the derivatives $\partial_{VV}P$, $\partial_{WW}P$ and a semi-center discretization of the cross derivative (see A.2), the numerical solution is obtained by inverting a matrix that is strongly diagonally dominant as long as the mesh-size does not change too sharply between two elements. The scheme will remain stable and positivity preserving with any time step. The inversion of the linear system is carried out by an LU pre–conditioned gradient procedure.

The one dimensional Fokker–Planck equations (6) are discretized by using the one-dimensional scheme from (Marpeau et al., 2009) (See A.4). The main difference here is the presence of the age variables, s and r. Since the age evolves simultaneously with time, we just solve Eq. (6) at each time step, regardless of age, and shift the age variable by one time step. Finally, Eq. (7) is solved exactly. The other boundary conditions given in Eqs. (8)–(14) are discretized as in one space dimension. The convergence criterion is satisfied as the residual of the numerical scheme decreases to a pre-defined value, 10^{-6} in our study.

3.2 Computing spike train statistics

The statistics of the number of threshold crossings of an IF model are of special interest. Using neuroscience terminology, we will refer to each threshold crossing as a *spike* and the sequence of threshold crossings as a *spike train*. Let the stochastic set functions $N_V(s,t)$ and $N_W(s,t)$ denote the number spikes during the time interval $[s,t]$ in cell V and W, respectively. The *instantaneous firing* rate of cell $X = V, W$ is defined as the instantaneous rate at which the corresponding population of cells spikes at time t. In terms of the spiking probability of a single cell, this can be written as

$$\nu_X(t) = \lim_{\Delta t \to 0} \frac{1}{\Delta t} \Pr(N_X(t, t+\Delta t) > 0).$$

The *conditional firing rate*, $\nu_{V|W}(\tau,t)$, is defined as the firing rate of cell V at time $t + \tau$ given that W has spiked at time t,

$$\nu_{V|W}(\tau,t) = \lim_{\Delta t \to 0} \frac{1}{\Delta t} \Pr\left(N_V(t+\tau+\Delta t, t+\tau) > 0 \,\middle|\, N_W(t, t+\Delta t) > 0\right)$$

and similarly for $\nu_{W|V}(\tau,t)$. The conditional firing rate can be normalized by the rates to obtain the spike train *cross-covariance function*

$$C_{VW}(\tau,t) = \nu_W(t)\left(\nu_{V|W}(\tau,t) - \nu_V(t+\tau)\right), \tag{18}$$

which is a common measure of correlation between the activity of two neurons over time. In the study of neural coding, it is often useful to know the propensity of one cell to spike during some time interval after another cell has spiked. For this purpose, we define the conditional mean rate,

$$S_{V|W}(a,b,t) = \frac{1}{(b-a)} \int_a^b \nu_{V|W}(t,\tau)d\tau \tag{19}$$

When the distribution of membrane potentials is in steady state, the spike trains are stationary and we can drop the explicit dependence on t to write ν_X, $\nu_{V|W}(\tau)$, $C_{VW}(\tau)$, and $S_{V|W}(a,b)$ without ambiguity.

Since action potentials are not explicitly modeled in the Fokker-Planck formalism described above, the spiking statistics must be calculated using properties of the probability density near threshold. The instantaneous firing rate of cell V can be obtained from the solution of the Fokker-Planck equation by taking the marginal flux over threshold,

$$\nu_V(t) = -D P_V(t,V)\big|_{V=V^T} \tag{20}$$

where $P_V(t,V) = \int_{W-\infty}^{W^T} P(t,V,W)dW$ is the marginal density of V. Thus, up to terms of $O((\Delta t)^2)$, the quantity $\nu_V(t)\Delta t$ is the probability mass that crossed threshold during the interval Δt. Equivalently, it equals the probability that a cell fires during this interval. The instantaneous firing rate of W is defined analogously.

The conditional firing rates are obtained by first calculating the conditional flux immediately after a spike in cell W at time t,

$$J_{\text{cond}}(t,V) := -D\,\partial_W P(t,V,W)\big|_{W=W^T}. \tag{21}$$

This conditional flux is then normalized to give the conditional density, $J_{\text{cond}}(t,V) \to J_{\text{cond}}(t,V) \left/ \int_{V-\infty}^{V^T} J_{\text{cond}}(t,x)dx \right.$ and used as an initial condition for the 1-dimensional Fokker-Planck equation,

$$\partial_\tau P_1(V,\tau) = -\partial_V(f(V) - D\,\partial_V P_1(V,\tau)). \tag{22}$$

As the solution of this equation evolves, the conditional firing rate of V is given by $\nu_{V|W}(\tau,t) = -D\,\partial_V P_1(V,\tau)$. The conditional firing rate can then be normalized, *cf.* Eq. (18), to get the cross-covariance function, or integrated, *cf.* Eq. (19), to get the conditional expected mean rate. We experienced convergence problems with the derivative of the finite volume solutions at the upper corner, (V^T, W^T), of the spatial domain. Due to the convergence issues discussed in Appendix A.6, the finite volume approximation to the conditional firing rate does not converge when τ is very small.
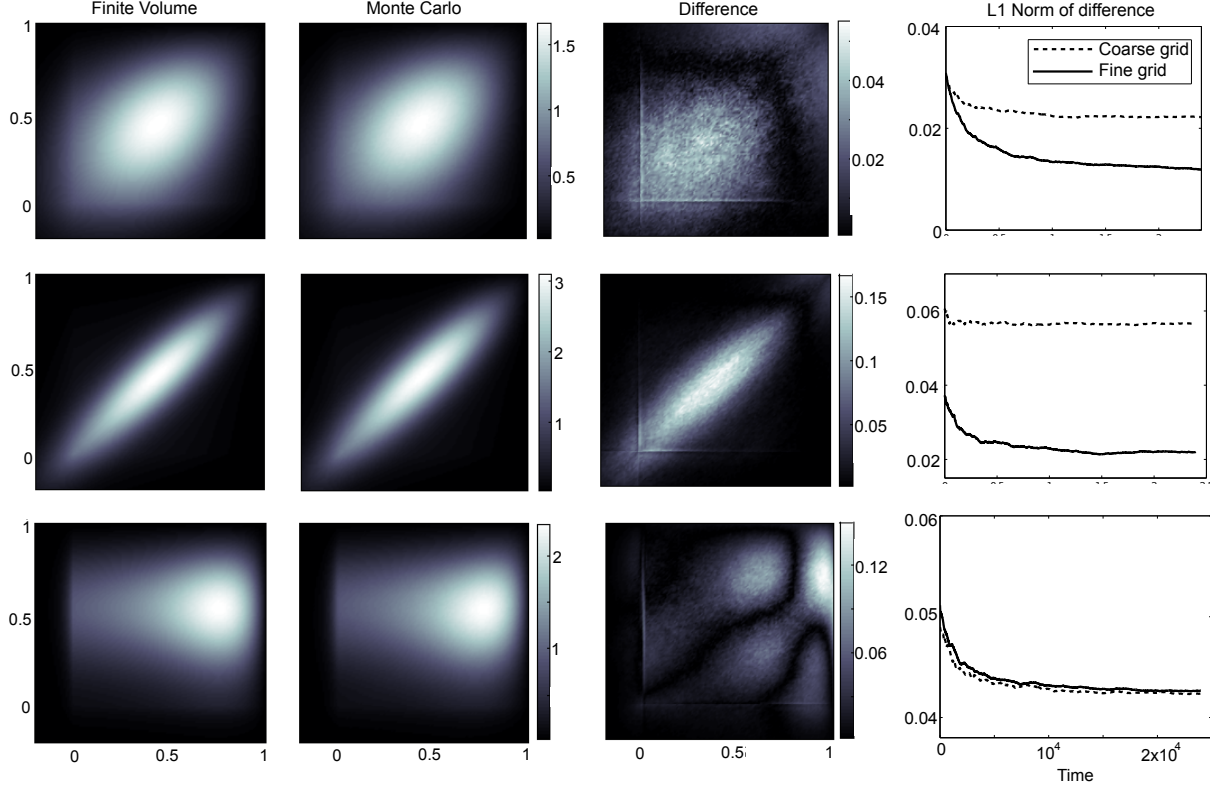
Figure 2: *Stationary probability densities for a pair of LIF neurons*: Results from finite volume simulations (first), MC simulations (second), the difference between the two approximations (third column). To test convergence of the finite volume method, we used a coarse (100×100 elements in the unit square), and a fine grid (200×200 elements). The fourth column shows the L^1 norm of the difference between the equilibrium distributions obtained using the finite volume and MC simulations. Parameters: for the top panel, $\mu_W = \mu_V = 0.5$, $D = 0.05$, $c = 0.5$, $\tau = 0.5$; for the middle panel, $\mu_W = \mu_V = 0.5$, $D = 0.05$, $c = 0.9$, $\tau = 0.5$; for the bottom panel, $\mu_V = 1.2, \mu_W = 0.6$, $D = 0.05$, $c = 0.3$, $\tau = 0.2$. The first three columns were obtained using a coarse (100×100) grid.

4   Validation of the numerical solution

As the finite volume numerical scheme we developed is novel, we first compare its output to that obtained using Monte Carlo (MC) simulations (See Appendix A.5). We consider both stationary and non-stationary inputs.

As an example we choose the case of two LIF neurons which corresponds to setting $f(V,W) = -V + \mu_V$, and $g(V,W) = -W + \mu_W$ in Eq. (3) (Burkitt, 2006). When $\mu_V < V^T$ and $\mu_W < W^T$ the cells are in the fluctuation dominated regime, and firing is due to large excursions of membrane voltages from the mean. As shown in top and middle panels of Fig. 2, the finite volume method provides an excellent approximation of the stationary distribution when the input to the two cells is constant in time. As the correlation between the inputs to the two cells, measured by c, increases, the membrane potentials become more correlated, and their joint probability density is stretched along the diagonal.

When $\mu_V > V^T$ or $\mu_W > W^T$ it is the DC component of the input current that drives the cells over threshold. This situation is somewhat more challenging to simulate, since much of the mass of the invariant distribution lies close to the threshold. The gradient of the solution close to the boundary becomes large. Together with the Dirichlet boundary conditions, this causes larger errors in the numerical approximation close to the boundary. The bottom panels of Fig. 2

7

show that the finite volume method still performs well in this situation.

We can change the drift term in Eq. (2) to simulate a different integrate–and–fire model. In particular, the quadratic integrate and fire (QIF) model is obtained by setting $f(V,W) = V^2 + \mu_V, g(V,W) = W^2 + \mu_W$ (Ermentrout and Kopell, 1986; Brunel and Latham, 2003). Fig. 3 demonstrates that the finite volume numerical scheme performs well in computing the invariant distribution for this model.

The finite volume scheme was designed to compute the evolution of the joint probability density of the two sub-threshold voltages in time. Stationary distributions were presented here for ease of visualization. A comparison of time dependent solutions obtained using finite volume and MC methods is available online at `http://www.math.uh.edu/~josic/myweb/research/papers/FV/`. The animation shows the time dependent density from $t = 100$ to $t = 110$ for a pair of LIFs with $\mu = |\sin(t)|$, $c = |\sin(t)|/2$, and $D = 0.1|\sin(t)|$.
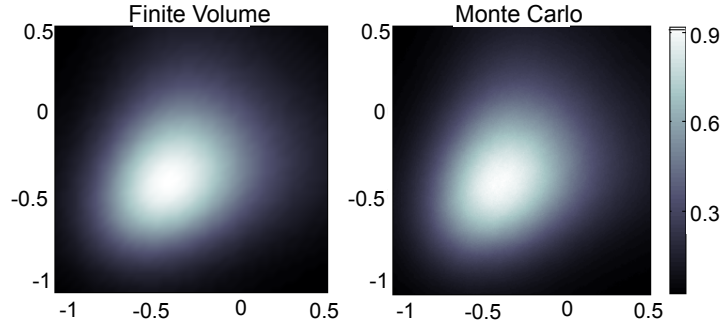

Figure 3: *Stationary probability densities for a pair of QIF neurons*: finite volume method (left), MC simulations (right) when $\mu_V = \mu_W = -0.1$, $D = 0.1$ and $c = 0.3$.

5 Gaussian approximation

The LIF model is ubiquitous in stochastic modeling of excitable systems primarily due to its mathematical tractability. Closed form expressions have been obtained for the stationary density, spiking statistics, and linear response properties of the one neuron model (Lindner, 2001). However, exact closed form expressions are not known for the two neuron model with $c \neq 0$ discussed in Section 2. The numerical methods we describe here offer a way of exploring the behavior of the LIF model in the absence of analytic solutions. However, even with fast numerical methods, exploring large regions of parameter space may not be possible. Approximate analytic solution, are therefore frequently necessary to gain a deeper understanding of the model.

Much recent work has focused on deriving such approximations using perturbative methods. Linear response theory was used to study the dependencies in the output of a call pair receiving correlated input (Lindner, 2001; Ostojić et al., 2009; de la Rocha et al., 2007; Shea-Brown et al., 2008). These solutions involve integrals that must be evaluated numerically. Simpler approximations can be obtained by ignoring the threshold and reset boundary conditions when the neurons are in the fluctuation dominated regime, and firing rates are low. Neurons in the cortex may reside in this regime under typical conditions (Ringach and Malone, 2007). Previous approximations obtained in this regime required smoothness assumptions on the trajectories of the membrane potentials (Burak et al., 2009; Tchumatchenko et al., 2010). Since solutions to Eq. (2) are nowhere differentiable when $D > 0$, a different approach must be used for the LIF driven by white noise inputs. We next describe such an approximation. (We note that a similar approach has been used to examine the response of integrate-and-fire neurons driven by filtered Gaussian noise (Badel et al., 2010).)

When firing rates are small, the boundary conditions have a small impact on the solution of Eq. (3) and an approximate solution can be obtained by solving the free boundary problem ($V^T, W^T \rightarrow \infty$). Since firing is rare, the amount of time spent in the refractory states is negligible and the refractory period can be ignored ($\tau = 0$). Under this approximation, the stochastic process $(V(t), W(t))$ is an Ornstein-Uhlenbeck process in $\mathbb{R}^2$ (Gardiner, 1985). Given

bivariate Gaussian initial conditions, the solution to the Fokker-Planck equation at any time is a bivariate Gaussian and can be computed in closed form. This Gaussian approximation is accurate when $X^T - \mu_X \gg \sqrt{2D}$ for $X = V, W$.

For simplicity, in this section we assume that the two neurons receive statistically identical inputs so that $\mu_V = \mu_W = \mu$. We further assume that the neurons are dynamically identical so that $g_V = g_W$, $V_{\text{rest}} = W_{\text{rest}}$, and $V^T = W^T$. The analysis is similar in the asymmetric case. Without loss of generality, we rescale space so that $V_{\text{rest}} = W_{\text{rest}} = 0$ and $V^T = W^T = 1$. To simplify calculations, we also time in units of the membrane time constants so that $g_V = g_W = 1$.

For instance, the marginal or conditional firing rates can be approximated by the flux of the time dependent Gaussian distribution over threshold. As shown in Appendix B, this flux can be written in terms of the mean (m) and the variance (σ^2) of the Gaussian, and the input diffusion coefficient (D) as

$$J(m, \sigma^2, D) := \frac{(1-m)D}{\sqrt{2\pi}\sigma^3} e^{\frac{-(\mu-1)^2}{2\sigma^2}}. \qquad (23)$$

The steady state firing rate, ν_∞ is obtained by taking $m = \mu$ and $\sigma^2 = D$ to get $\nu_\infty = J(\mu, D, D) = \frac{\alpha}{\sqrt{\pi}} e^{-\alpha^2}$ where $\alpha := (1-\mu)/\sqrt{2D}$. This expression can also be obtained using large deviation methods (Kampen, 2007; Lindner, 2001; Shea-Brown et al., 2008).

From an approximation of the conditional firing rate, the cross-covariance function can be obtained as (see Appendix B),

$$C_{V|W}(\tau) = \nu_\infty(H(\tau) - \nu_\infty) = \frac{1}{\pi}\alpha^2 e^{-\alpha^2} \left(\frac{e^{t - \frac{\alpha^2(e^\tau - c)}{c + e^\tau}}}{\sqrt{1 - c^2 e^{-2\tau}}(c + e^\tau)} - e^{-\alpha^2} \right).$$

In Fig. 4, we compare this approximation to the cross-covariance function to the cross-covariance function obtained from finite volume simulations. As expected, we find that the two agree well when firing rates and correlations are small, but disagree when μ, D or c are larger.

Further expressions for other stationary and non-stationary spiking statistics under the Gaussian approximation are derived in Appendix B. We use these approximations to examine the response of a pair of cells to time-varying inputs next.

5.1 Response to step changes in the input – single cell response

If a pair of cells responds rapidly to a change in an input parameter, then the output of the cell pair can accurately capture the information present in a time-varying input signal (Silberberg et al., 2004; Masuda, 2006). It is therefore useful to understand how the spiking statistics of a neuron or pair of neurons respond to changes in input parameters, μ, D, and c in Eq. (2). The response of the cell pair is measured by their joint firing rate (ν_V, ν_W), and we first examine how rapidly this response can track changes in the inputs to the model.

We start by examining the response of a single cell. Fig. 5 *Left* shows the time dependent firing rate after a step change in the mean, (μ, light line), and variance, (D, heavy line). We compare the response for the Gaussian approximation, derived in closed form in Appendix B, to the result from finite volume simulations. After a change in parameters, the distribution of (V, W), and therefore the spiking statistics, relax exponentially to a new steady state. However, the speed of this relaxation depends on the parameters that are changed. After a step change in the mean input, μ, the firing rate relaxes to a new steady state with a time constant of 1 (i.e., one membrane time constant). After a change in the variance, D, the firing rate jumps discontinuously, then approaches the new steady state value with a faster time constant of $1/2$ (see Appendix B). The fact that changes in the variance of the input are tracked faster than changes in input intensity is a fundamental property of the Ornstein-Uhlenbeck process (Gardiner, 1985). Therefore, changes in variance can be tracked more faithfully than changes in the DC component of the input. Related observations are made in (Khorsand and Chance, 2008; Hasegawa, 2009; Silberberg et al., 2004; Masuda, 2006), where mainly the discontinuous change in output firing rate in response to a step change in input variance was examined.

It also follows that a transient pulse change in D results in a larger transient in the firing rate than a comparable pulse change in μ. This prediction is verified for the Gaussian approximation and for finite volume simulations in Fig. 5 *Right*.
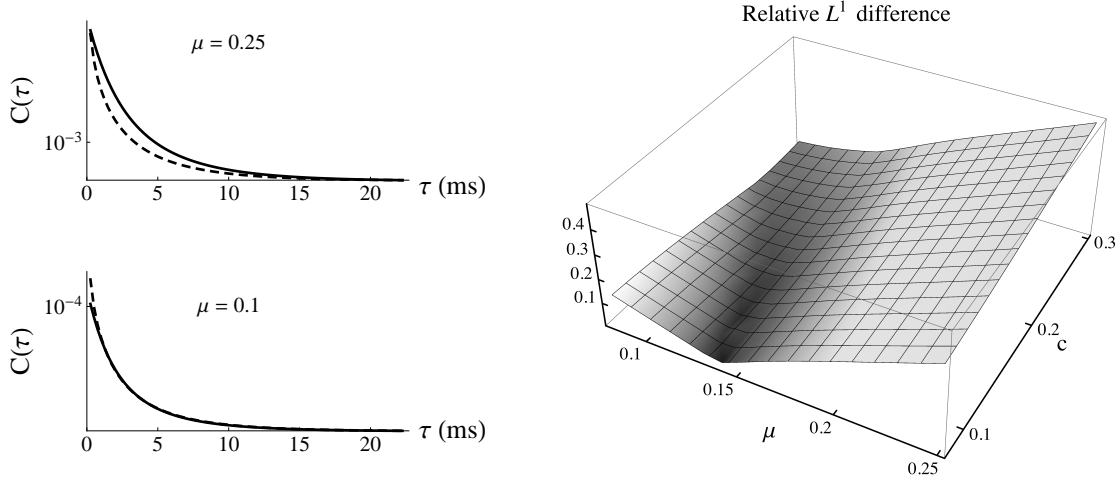
Figure 4: Left: Cross-covariance functions for $\mu = 0.1$ and $\mu = 0.25$ when $D = 0.05$ and $c = 0.2$. The solid lines were obtained from the Gaussian approximation and the dashed lines from finite volume simulations. Right: The relative L^1 difference between the Gaussian and finite volume cross-covariance functions (L^1 difference divided by the L^1 norm of the finite volume result) for $\mu \in [0.075, 0.25]$ and $c \in [0.075, 0.3]$. The L^1 norm was computed for $\tau > 0.15$, due to the convergence issues of $\nu_{V|W}$ for small τ discussed in Appendix A.6. The cross-covariance function has units Hz^2 In this figure and in Figs. 5 and 6, the axes are labeled assuming a membrane time constant of $1/g_V = 1/g_W = 5$ms.
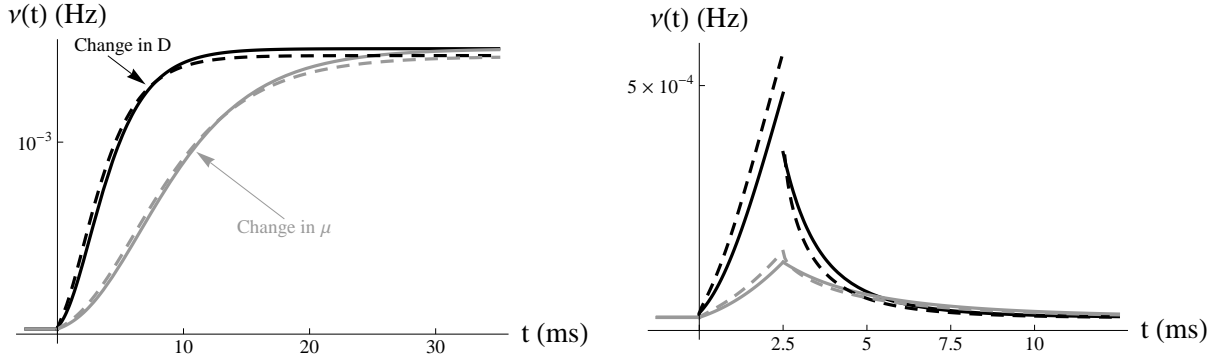

Figure 5: *Left*: Instantaneous firing rate after a step change in input statistics. Result from the Gaussian approximation are plotted as solid lines and finite volume simulations as dashed lines. For time $t < 0$, the parameters were set to $\mu = 0$ and $D = 0.03$. At time $t = 0$ the parameters were changed to $\mu = 0.1334$ (grey) or $D = 0.04$ (black). The values of μ and D were chosen so that the steady state firing rate after the change in parameters was the same whether μ or D was changed. *Right*: Instantaneous firing rate after a pulse change in input statistics. Same as Left, but the parameters were changed back to $\mu = 0$ and $D = 0.03$ at time $t = 2.5$ms. Note that a step change in the input variance, D, results in an instantaneous jump in the firing rate, followed by a continuous relaxation to the steady state. To illustrate the quantitative accuracy of the Gaussian approximation, we used parameters that resulted in low firing rates. Fig. 6 illustrates that, while the Gaussian approximation is less accurate when firing rates are moderate, the approximation can still capture the qualitative behavior of the spiking statistics.

10

5.2   Response to step changes in the input – joint response

We next examine how joint response of the cell pair in response to a step change in the input. The cross-covariance function, defined in Eq. (18), is commonly used to measure dependencies between two spike trains over time. Fig. 4 compares the Gaussian approximation of the cross-covariance function to that obtained using finite volume simulations. As expected, the two results agree when firing rates are low. However, as μ or D increase, the firing rate increases and the Gaussian approximation breaks down.

Fig. 6 *Top* shows the two-point conditional firing rate after a step change in the parameter D. This function completely characterizes the second order correlations of the two cells over time. Such a plot would be computationally prohibitively expensive to obtain using direct Monte-Carlo simulations, especially when firing rates are low. The parameters for Fig. 6 were chosen so that the Gaussian approximation does not agree quantitatively with the finite volume simulations. However, as shown in Fig. 6 *Bottom*, the Gaussian approximation successfully predicts the qualitative behavior of the bivariate spiking statistics with changing input parameters.

A pulse change in D has a larger impact on the propensity of the cells to fire together than a comparable pulse change in μ (compare to Fig. 5 *Right*). A pulse change in c has an intermediate impact. If a downstream cell that receives inputs from cells V and W is sensitive to synchrony in its inputs (Salinas and Sejnowski, 2000), then the cell would response more quickly and strongly to changes in D or c than to changes in μ.

This suggests that downstream cells that act as coincidence detectors are most sensitive to upstream changes in input variance, and respond more weakly to changes in input intensity.

6   Discussion

Population density methods have a long history in neuroscience. They have been used to study both the statistics of responses of single neuron (Tuckwell, 1988), and neural populations (Harrison, et al 2005). Studying the evolution of the ensembles, rather than tracking individual neurons has several advantages: While the dynamics of each individual cell is stochastic, their probability density evolves deterministically. The probability density is therefore easier to study analytically and numerically. For instance, both linear response methods (Lindner, 2001; de la Rocha et al., 2007; Ostojić et al., 2009), and the Gaussian approximation discussed here (see also (Burak et al., 2009; Tchumatchenko et al., 2010; Badel et al., 2010)), are obtained by considering the evolution of large populations in the diffusive limit. Simulating the evolution of population densities is typically orders of magnitudes faster than simulating the evolution of each individual cell in a population (Nykamp and Tranchina, 2000; Omurtag, 2000). For instance, obtaining the two–point time dependent firing rate, $\nu_{\text{cond}}(\tau, t)$ shown in Fig. 6, would not be feasible using Monte Carlo methods on an average machine today.

We have concentrated on a simple version of the model to keep the presentation relatively concise. For instance, it is easy to consider cells receiving different, potentially time dependent drives, $\mu_V(t), \mu_W(t), D_V(t)$ and $D_W(t)$. We have mainly considered drift terms of the form, $f(V,W) = f(V)$, and $g(W,V) = g(W)$, in Eq. (2), so that the two populations were uncoupled. Coupled populations have been considered earlier (Nykamp and Tranchina, 2000). We hence concentrated on examining the effects of anisotropic diffusion. However, the numerical methods we described can easily handle coupling between cells in the sub–threshold regime. This could be used to examine the interplay of correlated inputs and cell coupling (Schneider et al., 2006). However, we do not know whether there is a direct way to include super-threshold coupling in the present diffusive approximation. We also note that high firing rates can lead to steep gradients of the probability density close to the boundary and convergence problems in the numerical methods we used. This suggests that numerical techniques will have to be developed further to accurately capture the response of IF neurons in this regime.

We have used our numerical and analytical approach to examine the best way to transmit information in a pair of cells. We found that both the single cell response and the joint response tracks changes in noise intensity more accurately than changes in the mean drive or correlations in the cell inputs. Thus input variance appears to provide the best channel to code information at the single cell and population level.

The numerical methods we have developed can be used to further examine how the output of a cell pair reflects their interactions and dependencies in their inputs. It is of particular interest how cells respond to signals that vary in
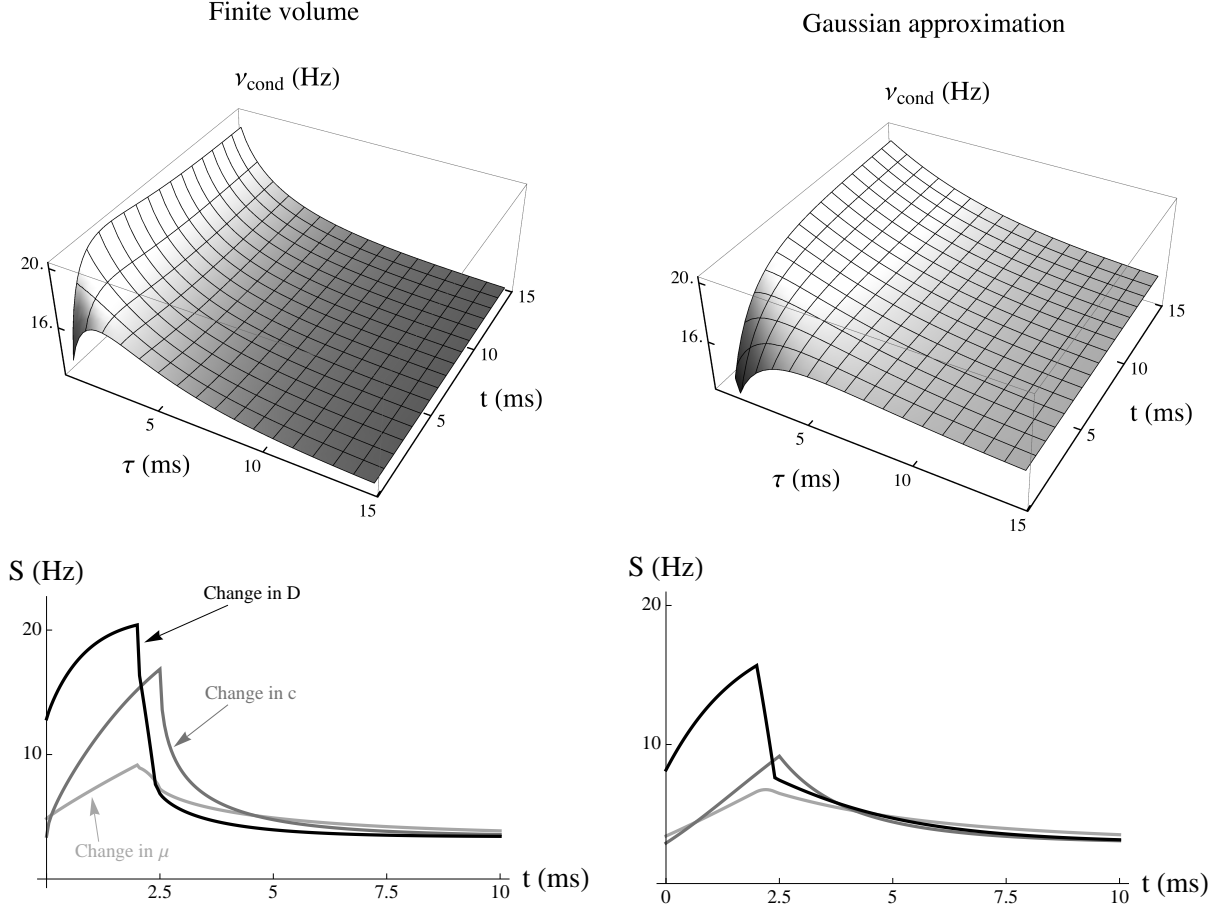
Figure 6: *Top row*: The two-point time dependent firing rate, $\nu_{\text{cond}}(\tau,t)$ after a step change in D with $D = 0.1$ for $t < 0$ and $D = 0.2$ for $t > 0$. Parameters $\mu = 0$ and $c = 0.1$ were held constant. Values for $\tau \leq 0.5$ms were omitted due to the convergence issues discussed in Appendix A.6. *Bottom row*: The conditional mean rate, S, after a pulse change in parameters. The parameters changed from $D, c, \mu = 0.1, 0.1, 0$ for $t < 0$ to $D = 0.2$ (black), $c = 0.435$ (dark grey), or $\mu = 0.293$ (light grey) for $t \in [0, 2.5]$ms, then back to $D, c, \mu = 0.1, 0.1, 0$ for $t > 2.5$ms. The conditional mean rate, calculated according to Eq. (19) with $a = 0.15$ms and $b = 0.5$ms, measures the propensity for cell V to spike within the first 0.5ms after cell W spikes. We chose $a = 0.15$ms to circumvent the convergence issues discussed in Appendix A.6. *Left column*: finite volume simulations. *Right column*: Gaussian approximation.
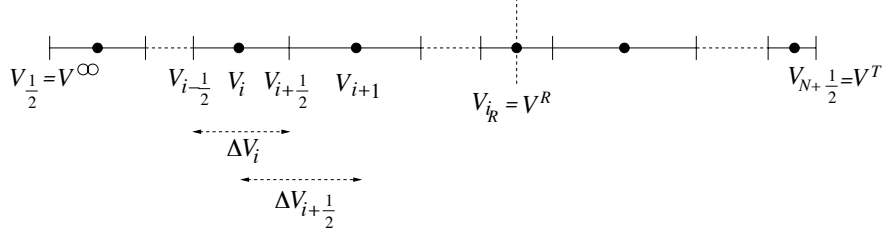
12

Figure 7: A schematic depiction of the subdivision of the domain $(V^{-\infty}, V^T)$ into subintervals $(V_{i-\frac{1}{2}}, V_{i+\frac{1}{2}})$. The reset voltage V^R is at V_{i_R}. Similar notations are used for the second neuron voltage W.

time. The finite volume method we described is well suited to this task, as it is designed to capture the time–dependent response of a cell pair.

A   Numerical schemes

In the following we provide a description of the finite volume method used in the numerical simulations. Some of the details of the implementation are not standard. As we are not aware of a similar treatment of this type of equation, we give a detailed discussion of the novel aspects of the algorithm.

The time steps are defined by $\Delta t_n = t_{n+1} - t_n$. When there is no ambiguity, the time step is denoted by Δt. The intervals $(V^{-\infty}, V^T)$ and $(W^{-\infty}, W^T)$ are partitioned into N_V and N_W sub-intervals respectively. We denote ΔV_i as the ith step in the V-direction and ΔW_j as the jth step in the W-direction, for $i = 1, \ldots, N_V$, $j = 1, \ldots, N_W$. Our quadrilateral mesh elements, $Q_{i,j}$, are then defined by

$$Q_{i,j} = (V_{i-\frac{1}{2}}, V_{i+\frac{1}{2}}) \times (W_{j-\frac{1}{2}}, W_{j+\frac{1}{2}}),$$

with $V_{i-\frac{1}{2}} = \sum_{l=1}^{i-1} \Delta V_l$, $W_{j-\frac{1}{2}} = \sum_{l=1}^{j-1} \Delta W_l$. Our mesh-points (V_i, W_j) are the centroids of the cells, thus $V_i = V_{i-\frac{1}{2}} + \frac{1}{2}\Delta V_i$, $W_j = W_{j-\frac{1}{2}} + \frac{1}{2}\Delta W_j$. We make sure that there exist two indices i_R and j_R such that $V_{i_R} = V^R$ and $W_{i_R} = W^R$, which means that the two reset potentials fall exactly on some mesh-points, see Fig. 7. For every function ξ defined on $(0,T) \times \Omega$, the notation $\xi_{\alpha,\beta}^n$ stands for the approximation of $\xi(t_n, (V_\alpha, W_\beta))$, for $\alpha = i, i \pm \frac{1}{2}$, $\beta = j, j \pm \frac{1}{2}$. We denote ξ^n as the sequence $\{\xi_{i,j}^n\}_{i,j}$.

A.1   Treatment of the drift operator: scheme $\mathscr{S}_1$

The advection equation in (15) is discretized by using the one-dimensional numerical fluxes in (Marpeau et al., 2009) in each direction. We set

$$P_{i,j}^{n+1} = P_{i,j}^n - \frac{\Delta t}{\Delta v_i}\left(\mathscr{A}_{i+\frac{1}{2},j}^n - \mathscr{A}_{i-\frac{1}{2},j}^n\right) - \frac{\Delta t}{\Delta w_j}\left(\mathscr{A}_{i,j+\frac{1}{2}}^n - \mathscr{A}_{i,j-\frac{1}{2}}^n\right), \tag{24}$$

where the numerical fluxes are defined by

$$\mathscr{A}_{i+\frac{1}{2},j}^n = f_{i+\frac{1}{2},j}^+ P_{i,j}^n + f_{i+\frac{1}{2},j}^- P_{i+1,j}^n$$
$$+ \frac{1}{2}\frac{\Delta P_{i+\frac{1}{2},j}^n}{\Delta v_{i+\frac{1}{2}}}\left(f_{i+\frac{1}{2},j}^+ \Delta v_i \varphi\left(r_{i+\frac{1}{2},j}^p, \frac{\Delta v_{i+\frac{1}{2}}}{\Delta v_i}\right) - f_{i+\frac{1}{2},j}^- \Delta v_{i+1} \varphi\left(r_{i+\frac{1}{2},j}^m, \frac{\Delta v_{i+\frac{1}{2}}}{\Delta v_{i+1}}\right)\right), \tag{25}$$

13

for $1 \leq i \leq N_V - 1$, $1 \leq j \leq N_W$ and

$$
\mathscr{A}^n_{i,j+\frac{1}{2}} = g^+_{i,j+\frac{1}{2}} P^n_{i,j} + g^-_{i,j+\frac{1}{2}} P^n_{i,j+1}
$$
$$
+ \frac{1}{2} \frac{\Delta P^n_{i,j+\frac{1}{2}}}{\Delta w_{i,j+\frac{1}{2}}} \left(g^+_{i,j+\frac{1}{2}} \Delta w_j \varphi\left(r^p_{i,j+\frac{1}{2}}, \frac{\Delta w_{j+\frac{1}{2}}}{\Delta w_j}\right) - g^-_{i,j+\frac{1}{2}} \Delta w_{j+1} \varphi\left(r^m_{i,j+\frac{1}{2}}, \frac{\Delta w_{j+\frac{1}{2}}}{\Delta w_{j+1}}\right)\right), \quad (26)
$$

for $1 \leq i \leq N_V$, $1 \leq j \leq N_W - 1$, with notations $\Delta P^n_{i+\frac{1}{2},j} := P^n_{i+1,j} - P^n_{i,j}$, $\Delta P^n_{i,j+\frac{1}{2}} := P^n_{i,j+1} - P^n_{i,j}$,

$$
r^p_{i+\frac{1}{2},j} = \frac{f^+_{i-\frac{1}{2},j} \Delta P_{i-\frac{1}{2},j}}{f^+_{i+\frac{1}{2},j} \Delta P_{i+\frac{1}{2},j}}, \quad r^m_{i-\frac{1}{2},j} = \frac{f^-_{i+\frac{1}{2},j} \Delta P_{i+\frac{1}{2},j}}{f^-_{i-\frac{1}{2},j} \Delta P_{i-\frac{1}{2},j}}, \quad r^p_{i,j+\frac{1}{2}} = \frac{g^+_{i,j-\frac{1}{2}} \Delta P_{i,j-\frac{1}{2}}}{g^+_{i,j+\frac{1}{2}} \Delta P_{i,j+\frac{1}{2}}}, \quad r^m_{i,j-\frac{1}{2}} = \frac{g^-_{i,j+\frac{1}{2}} \Delta P_{i,j+\frac{1}{2}}}{g^-_{i,j-\frac{1}{2}} \Delta P_{i,j-\frac{1}{2}}} \ ,
$$

where the limiter function φ is defined by $\varphi(a,b) = 2b \max\big(0, \min(1,2a), \min(a,2)\big)$.

To comply with Dirichlet boundary conditions in (15), we further impose $\mathscr{A}^n_{0,j} = \mathscr{A}^n_{N_V+\frac{1}{2},j} = \mathscr{A}^n_{i,0} = \mathscr{A}^n_{i,N_W+\frac{1}{2}} = 0$, for $1 \leq i \leq N_V$, $1 \leq j \leq N_W$.

Proposition 1 *The numerical scheme (24) is non-negativity preserving under the* CFL *condition*

$$
\Delta t \left(\frac{f^+_{i-\frac{1}{2},j} - f^-_{i+\frac{1}{2},j}}{\Delta v_i} + \frac{g^+_{i,j-\frac{1}{2}} - g^-_{i,j+\frac{1}{2}}}{\Delta w_j} + \left(\frac{f_{i+\frac{1}{2},j} - f_{i-\frac{1}{2},j}}{\Delta v_i}\right)^+ + \left(\frac{g_{i,j+\frac{1}{2}} - g_{i,j-\frac{1}{2}}}{\Delta w_j}\right)^+ \right) \leq 1 \ .
$$

The proof of this proposition is similar to the proof of the corresponding one-dimensional result in (Marpeau et al., 2009), and we therefore omit it here.

A.2 Treatment of the diffusion operator: scheme $\mathscr{S}_2$

Our approximation of the solutions to (16) in two space dimensions is given by the implicit scheme

$$
\frac{\Delta v_i \Delta w_j}{\Delta t} P^{n+1}_{i,j} - \left(\mathscr{B}^{n+1}_{i+\frac{1}{2},j} - \mathscr{B}^{n+1}_{i-\frac{1}{2},j}\right) - \left(\mathscr{B}^{n+1}_{i,j+\frac{1}{2}} - \mathscr{B}^{n+1}_{i,j-\frac{1}{2}}\right) = \frac{\Delta v_i \Delta w_j}{\Delta t} P^n_{i,j} + \delta_{i,i_R} S^{W,n+1}_j + \delta_{j,j_R} S^{V,n+1}_i, \quad (27)
$$

where δ_{k_1,k_2} is the symbol of Kronecker and $S^{V,n+1}_i$ and $S^{W,n+1}_j$ account for the re-injection condition (11), see subsection A.3. The numerical diffusive fluxes are defined by

$$
\mathscr{B}^{n+1}_{i+\frac{1}{2},j} = D \frac{P^{n+1}_{i+1,j} - P^{n+1}_{i,j}}{\Delta v_{i+\frac{1}{2}}} + \frac{CD}{2} \left(\frac{P^{n+1}_{i+1,j+1} - P^{n+1}_{i+1,j}}{\Delta w_{j+\frac{1}{2}}} + \frac{P^{n+1}_{i,j} - P^{n+1}_{i,j-1}}{\Delta w_{j-\frac{1}{2}}} \right), \quad (28)
$$

for $1 \leq i \leq N_V - 1$, $1 \leq j \leq N_W$ and

$$
\mathscr{B}^{n+1}_{i,j+\frac{1}{2}} = D \frac{P^{n+1}_{i,j+1} - P^{n+1}_{i,j}}{\Delta w_{j+\frac{1}{2}}} + \frac{CD}{2} \left(\frac{P^{n+1}_{i+1,j+1} - P^{n+1}_{i,j+1}}{\Delta v_{i+\frac{1}{2}}} + \frac{P^{n+1}_{i,j} - P^{n+1}_{i-1,j}}{\Delta v_{i-\frac{1}{2}}} \right), \quad (29)
$$

for $1 \leq i \leq N_V$, $1 \leq j \leq N_W - 1$.

Remark 1 *Notice that the second term of the right-hand side in (28) stand for a centered finite difference discretization of the cross derivative $CD\partial^2_{vw}P$ on the right vertical interface of $Q_{i,j}$. Other numerical schemes have been implemented in (Bruneau et al., 2005; Rasetarinera, 1995; Bourgeat and Kern, 2004), but yield unconditionally positive off-diagonal coefficients in the diffusion matrix, therefore producing negative undershoot near sharp solution gradients. When the neurons are strongly correlated (i.e.: when $C \approx 1$), the gradients of the solution can be very sharp. The advantage of our method is that all the off-diagonal coefficients are nonnegative where the mesh is uniform, which means, the region where the solution is not 0 in practice. Using a similar remark in (29), the resulting numerical scheme (27) is nonnegative in realistic applications.*
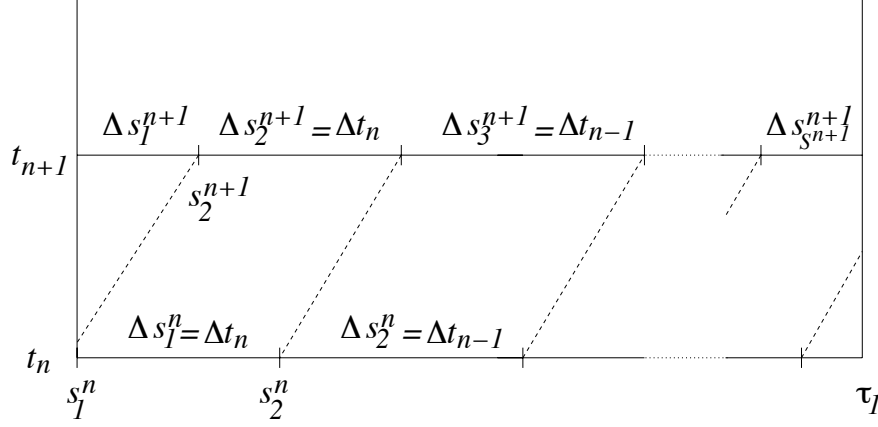
14

Figure 8: A schematic depiction of the time dependent subdivision of the domain $(0, \tau_V)$ into subintervals (s_k^n, s_{k+1}^n).

The boundary conditions in (16) are implemented as $\mathscr{B}_{\frac{1}{2},j}^{n+1} = \mathscr{B}_{i,\frac{1}{2}}^{n+1} = 0$ and

$$\mathscr{B}_{N_V+\frac{1}{2},j}^{n+1} = D\frac{0 - P_{N_V,j}^{n+1}}{\Delta V_{N_V}/2} \ , \ \mathscr{B}_{i,N_W+\frac{1}{2}}^{n+1} = D\frac{0 - P_{i,N_W}^{n+1}}{\Delta W_{N_W+\frac{1}{2}}} \ ,$$

$$\mathscr{B}_{i+\frac{1}{2},N_W}^{n+1} = D\frac{P_{i+1,N_W}^{n+1} - P_{i,N_W}^{n+1}}{\Delta V_{i+\frac{1}{2}}} + \frac{CD}{2}\left(\frac{0 - P_{i+1,N_W}^{n+1}}{\Delta W_{N_W}/2} + \frac{P_{i,N_W}^{n+1} - P_{i,N_W-1}^{n+1}}{\Delta W_{N_W-\frac{1}{2}}}\right) \ ,$$

$$\mathscr{B}_{N_V,j+\frac{1}{2}}^{n+1} = D\frac{P_{N_V,j+1}^{n+1} - P_{N_V,j}^{n+1}}{\Delta W_{j+\frac{1}{2}}} + \frac{CD}{2}\left(\frac{0 - P_{N_V,j+1}^{n+1}}{\Delta V_{N_V}/2} + \frac{P_{N_V,j}^{n+1} - P_{N_V-1,j}^{n+1}}{\Delta V_{N_V-\frac{1}{2}}}\right) \ .$$

A.3 Treatment of the re-injection condition (11)

Since the time variable t and the age variable $s \in (0, \tau_1)$ evolve together, the domain $(0, \tau_1)$ is dynamically partitioned into sub-intervals (s_k^n, s_{k+1}^n) such that $s_k^n = 0$ and $s_{k+1}^n = \min(s_k^n + \Delta s_k^n, \tau_1)$, where the age steps Δs_k^n match the time steps as follows: for all n, $\Delta s_1^n = \Delta t_n$ and $\Delta s_k^n = \Delta t_{n-k}$, $k = 1, \ldots, n-1$. We set $K_n^s = \max\{k, \ s_k^n < \tau_1\}$, see figure 8. In the same way, we discretize $(0, \tau_2)$ into sub-intervals (r_k^n, r_{k+1}^n) such that $r_k^n = 0$ and $r_{k+1}^n = \min(r_k^n + \Delta r_k^n, \tau_1)$, with $\Delta r_1^n = \Delta t_n$, $\Delta r_k^n = \Delta t_{n-k}$, $k = 1, \ldots, n-1$. We set $K_n^r = \max\{k, \ r_k^n < \tau_2\}$.

Then, defining the piecewise constant functions

$$\tilde{R}_{1,i}(t) = \sum_{n=0}^{+\infty} R_{1,i}^{n+1,K_n^r}\chi_{(t_n,t_{n+1}]}(t) \ , \ \tilde{R}_{2,j}(t) = \sum_{n=0}^{+\infty} R_{2,j}^{n+1,K_n^s}\chi_{(t_n,t_{n+1}]}(t) \ ,$$

the quantities

$$S_i^V := \int_{\tau_2-\Delta t_n}^{\tau_2} \tilde{R}_{1,i}(t)\,dt \text{ and } S_j^W := \int_{\tau_1-\Delta t_n}^{\tau_1} \tilde{R}_{2,j}(t)\,dt$$

are injected into (27).

A.4 Treatment of the one neuron boundary condition (6)–(10)

Equation (6) is discretized with the one–dimensional numerical scheme in (Marpeau et al., 2009). We use the operator splitting technique

$$R_{1,i}^{n+\frac{1}{2},k+\frac{1}{2}} = R_{1,i}^{n,k} - \frac{\Delta t}{\Delta V_i}(\mathscr{A}_{i+\frac{1}{2}}^n - \mathscr{A}_{i-\frac{1}{2}}^n) \ , \text{ and } R_{1,i}^{n+1,k+1} - \frac{\Delta t}{\Delta V_i}(\mathscr{B}_{i+\frac{1}{2}}^{n+1} - \mathscr{B}_{i-\frac{1}{2}}^{n+1}) = R_{1,i}^{n+\frac{1}{2},k+\frac{1}{2}} + \delta_{i,i_R}S^{W,n+1}, \quad (30)$$

15

where the advective and diffusive numerical fluxes are

$$\mathscr{A}^n_{i+\frac{1}{2}} = f^+_{i+\frac{1}{2},N_W} R^{n,k}_{1,i} + f^-_{i+\frac{1}{2},N_W} R^{n,k}_{1,i+1} + \frac{1}{2} \frac{\Delta R^{n,k}_{1,i+\frac{1}{2}}}{\Delta V_{i+\frac{1}{2}}} \left(f^+_{i+\frac{1}{2},N_W} \Delta V_i \varphi \left(r^p_{i+\frac{1}{2}}, \frac{\Delta v_{i+\frac{1}{2}}}{\Delta v_i} \right) - f^-_{i+\frac{1}{2},N_W} \Delta V_{i+1} \varphi \left(r^m_{i+\frac{1}{2}}, \frac{\Delta v_{i+\frac{1}{2}}}{\Delta v_{i+1}} \right) \right),$$

$$\mathscr{B}^{n+1}_{i+\frac{1}{2}} = D \frac{R^{n+1,k+1}_{1,i+1} - R^{n+1,k+1}_{1,i}}{\Delta v_{i+\frac{1}{2}}} \tag{31}$$

for $i = 1, N_V - 1$, and $\mathscr{A}^n_{\frac{1}{2}} = \mathscr{A}^n_{N_V+\frac{1}{2}} = \mathscr{B}^{n+1}_{\frac{1}{2}} = 0$, $\mathscr{B}^{n+1}_{N_V+\frac{1}{2}} = -\frac{D}{2\Delta V_{N_V}} R^{n+1,k+1}_{1,N_V}$.

We use similar arguments to discretize Eq. (6). Then, since the age steps Δs^n_k match the time steps Δt_{n-k}, equation (7) is solved exactly as $R^{n+1,k+1,l+1} = R^{n,k,l}$, and its boundary conditions (9) translate into $R^{n,0,l} = -\frac{D}{2\Delta V_{N_V}} R^{n,l}_{1,N_V}$, $R^{n,k,0} = -\frac{D}{2\Delta W_{N_W}} R^{n,k}_{2,N_W}$.

Finally, the first re–injection condition in (10) is taken into account by setting $S^{W,n+1,l} := R^{n+1,K^s,l}$ in (30). The second re–injection condition in (10) is discretized in a similar fashion.

A.5 Comparison with solutions obtained using Monte Carlo methods

Monte Carlo simulations were performed by integrating the Langevin equations (2) using the Euler–Maruyama method with time step 10^{-3} (Kloeden and Platen, 1992). Histograms in the $V - W$ plane were typically created using 2×10^7 points sampled from a long, simulated trajectory and using the same grid on the domain $\Omega = (V^{-\infty}, V^T) \times (W^{-\infty}, W^T)$ as in the finite volume simulation. To compare the results of the finite volume and Monte Carlo methods, we computed the L^1 norm of their difference by taking the absolute value of the difference at each grid point and summing over the domain Ω.

A.6 Calculating spike train statistics from finite volume simulations

To obtain spike train statistics from finite volume simulations, we used the definitions from Sec. 3.2. We solved the two-dimensional Fokker-Planck equation using the numerical scheme described above. To solve the one-dimensional Fokker-Planck equation (for example, to obtain instantaneous or conditional firing rates $cf.$ Eq. (22)), we used the one-dimensional scheme described in (Marpeau et al., 2009) or we used the two-dimensional scheme and calculated the marginal density from the two-dimensional density as $P_1(t,V) = \int_{W^{-\infty}}^{W^T} P(t,V,w) dw$.

However, we met with convergence problems in calculating the conditional flux $J_{\text{cond}}(t,V)$ using Eq. (21). These are due to the fact that when W is close to the threshold value W^T, and the two neurons fire nearly synchronously, the re-injection process in the V and W directions is relatively complicated. Given these convergence issues, we ignored the first 0.5ms in the left panel of Fig. 4 and 0.75ms in right panel of Fig. 4 and Fig. 6 when we computed the conditional firing rate $\nu_{V|W}(\tau,t)$.

B   Gaussian approximation of the LIF

In this section we derive an approximation of the spiking statistics for the LIF in low firing rate regimes. Recall that the LIF is defined by taking $f(V,W) = -g_V(V - V_{\text{rest}})$ and $g(V,W) = -g_W(W - W_{\text{rest}})$ in Eq. (2). For simplicity, in this section we assume that the two neurons receive statistically identical inputs so that $\mu_V = \mu_W = \mu$. We further assume that the neurons are dynamically identical so that $g_V = g_W$, $V_{\text{rest}} = W_{\text{rest}}$, and $V^T = W^T$. The analysis is similar in the asymmetric case. Without loss of generality, we rescale space so that $V_{\text{rest}} = W_{\text{rest}} = 0$ and $V^T = W^T = 1$. To simplify calculations, we also time in units of the membrane time constants so that $g_V = g_W = 1$.

When $\alpha := (1-\mu)/\sqrt{2D}$ is large, firing rates are low and the boundary conditions at threshold have a small impact on the distribution. Also, since the cells only spike rarely, the refractory period can be ignored. In such regimes, the solution of the full problem is approximated by the solution of the free boundary problem. This approximation, which

we call the Gaussian approximation, is accurate in the limit $\alpha \to \infty$. In this case the membrane potentials $(V(t), W(t))$ are described by an Ornstein-Uhlenbeck process on $\mathbb{R}^2$. Such processes are well-understood and the spiking statistics can be computed exactly, as we show below.

Assume that the initial distribution $P(0, U)$ is a bivariate Gaussian with marginal means $m(0) = E[V(0)] = E[W(0)]$, variance $\sigma^2(0) = \text{var}(V(0)) = \text{var}(W(0))$, and covariance $\gamma(0) = \text{cov}(V(0), W(0))$. Then the solution at any time $t \geq 0$ (in the absence of boundary conditions) is Gaussian with mean, variance and covariance given respectively by

$$
\begin{aligned}
m(t) &= e^{-t}m(0) + (1 - e^{-t})m(\infty), \\
\sigma^2(t) &= e^{-2t}\sigma^2(0) + (1 - e^{-2t})\sigma^2(\infty), \text{ and} \\
\gamma(t) &= e^{-2t}\gamma(0) + (1 - e^{-2t})\gamma(\infty)
\end{aligned}
$$

where

$$
m(\infty) = \mu, \quad \sigma^2(\infty) = D, \quad \text{and} \quad \gamma(\infty) = cD
$$

are the steady state mean, variance, and covariance.

B.1 Conditional firing rate and spike count correlation in the steady state

The results above can be used to derive an approximation of the steady state conditional firing rates. Since the joint distribution of (V, W) is a bivariate Gaussian, the distribution of V given that $W = W^T = 1$ is a univariate Gaussian. The conditional mean and variance are $m_c(0) = c(1 - \mu) + \mu$ and $\sigma_c^2(0) = D(1 - c^2)$ respectively. As time evolves, the conditional density of V relaxes to its steady state. The density during this relaxation is a univariate Gaussian with mean

$$
m_c(\tau) = e^{-\tau}m_c(0) + (1 - e^{-\tau})m_c(\infty)
$$

and variance

$$
\sigma_c^2(\tau) = e^{-2\tau}\sigma_c^2(0) + (1 - e^{-2\tau})\sigma_c^2(\infty)
$$

where $m_c(\infty) = \mu$ and $\sigma_c^2(\infty) = D$ are the stationary mean and variance. Note that for c near 1, $m_c(0)$ is near $V^T = 1$ which violates the assumptions of the Gaussian approximation, namely that the mass near threshold is small. In this case, the *conditional* flux across threshold is large even if the *marginal* fluxes across threshold are small. Thus, for the approximation of the conditional firing rate to be accurate, we must assume that α is large *and* that c is small.

The conditional firing rate is simply

$$
\nu_{V|W}(\tau) = J(\mu_c(\tau), \sigma_c^2(\tau), D)
$$

where $J(\mu, \sigma^2, D)$ is as defined in (23). This expression does not depend on t because we have assumed that the two-dimensional distribution is in its steady state. Later, we look at the two point conditional firing rate outside of the steady state.

The steady state cross-covariance function is obtained from the conditional firing rate *cf.* Eq. (18) to obtain

$$
C_{V|W}(\tau) = \nu_\infty(H(\tau) - \nu_\infty) = \frac{1}{\pi}\alpha^2 e^{-\alpha^2} \left(\frac{e^{t - \frac{\alpha^2(e^\tau - c)}{c + e^\tau}}}{\sqrt{1 - c^2 e^{-2\tau}}(c + e^\tau)} - e^{-\alpha^2} \right)
$$

To first order in c this gives,

$$
R(\tau) = \frac{c}{\pi}\alpha^2(2\alpha^2 - 1)e^{-2\alpha^2 - \tau} + o(c^2). \tag{32}
$$

The asymptotic spike count correlation, defined by

$$
\rho := \lim_{t \to \infty} \frac{\text{cov}(N_V(t), N_W(t))}{\sqrt{\text{var}(N_V(t))\text{var}(N_W(t))}},
$$

17

can be written in terms of the conditional firing rate as (Shea-Brown et al., 2008)

$$\rho = \frac{2 \int_0^{\infty} \left(\nu_{V|W}(\tau) - \nu_{\infty} \right) d\tau}{CV^2}$$

where CV is the coefficient of variation of the spike train inter-spike intervals. In the low firing rate, $\alpha \to \infty$ limit, $CV \to 1$ and therefore, to first order in c and ν_{∞},

$$
\begin{aligned}
\rho &\approx 2 \int_0^{\infty} \left(\nu_{V|W}(\tau) - \nu_{\infty} \right) d\tau \\
&= \frac{c}{\sqrt{\pi}} 2\alpha \left(2\alpha^2 - 1 \right) e^{-\alpha^2}.
\end{aligned}
\tag{33}
$$

Since both ν_{∞} and the correlation susceptibility, $T := \frac{\rho}{c}$, are functions of the single parameter α and since ν_{∞} is monotonic with α, we may conclude that T is a function of ν_{∞} to first order in c and ν_{∞}. This same conclusion was reached in Shea-Brown et al. (2008) using linear response theory, though the expression derived for ρ,

$$\rho \approx \frac{c}{\sqrt{\pi}} \alpha \left(2\alpha - \frac{1}{\alpha} \right)^2, e^{-\alpha^2}
\tag{34}$$

differs from Eq. (33). For both expressions, $\frac{\partial \rho}{\partial \nu_{\infty}} \sim 4c\alpha^2$ as $\alpha \to \infty$ (i.e., as $\nu_{\infty} \to 0$). Comparing these two approximations to a more accurate linear response approximation (also from (Shea-Brown et al., 2008)), we found that Eq. (34) is more accurate than Eq. (33).

B.2   Time dependent input statistics

We will now investigate how the spiking statistics track time dependent changes in the inputs. When the input parameters to the neurons are time dependent, the two dimensional density $p(t,V,W)$ at any time t is a bivariate Gaussian whenever the initial condition is a bivariate Gaussian. Thus we can use the same methods as above to derive the time dependent spiking statistics. To illustrate the effects of time-dependent inputs, we concentrate on a simple time-dependent input model. We assume that each cell receives input with mean μ_0, diffusion D_0, and correlation c_0 for $t < 0$ and, at time $t = 0$, the input parameters change instantaneously to μ_1, D_1, and c_1. At some later time $t_0 > 0$, the inputs change back to the original values, μ_0, D_0, and c_0. A small value of t_0 models a pulse change in the inputs. Taking $t_0 = \infty$ models a step change. The discussion here can easily be generalized to arbitrary time-dependent input (e.g., sinusoidally varying inputs) by solving a simple linear ODE for the time dependent mean, variance, and covariance (Gardiner, 1985).

We assume that for time $t \leq 0$, the distribution is in its steady state so that

$$
\left.
\begin{aligned}
m(t) &= \mu_0, \\
\sigma^2(t) &= D_0 \\
\gamma(t) &= cD_0
\end{aligned}
\right\} \quad t \leq 0
$$

At time $t = 0$, the input statistics change and the mean and covariance matrix begin to track this change. In particular,

$$
\left.
\begin{aligned}
m(t) &= e^{-t}\mu_0 + \left(1 - e^{-t}\right)\mu_1, \\
\sigma^2(t) &= e^{-2t}D_0 + \left(1 - e^{-2t}\right)D_1 \\
\gamma(t) &= e^{-2t}c_0 D_0 + \left(1 - e^{-2t}\right)c_1 D_1
\end{aligned}
\right\} \quad t \in [0, t_0]
$$

At time t_0, the input statistics change back to μ_0, D_0, and c_0 and the distribution relaxes back to its steady state. In particular,

$$
\left.
\begin{aligned}
m(t) &= e^{-(t-t_0)}m(t_0) + \left(1 - e^{-(t-t_0)}\right)\mu_0, \\
\sigma^2(t) &= e^{-2(t-t_0)}\sigma^2(t_0) + \left(1 - e^{-2(t-t_0)}\right)D_0 \\
\gamma(t) &= e^{-2(t-t_0)}\gamma(t_0) + \left(1 - e^{-2(t-t_0)}\right)c_0 D_0
\end{aligned}
\right\} \quad t \geq t_0
$$

where $\mu(t_0)$, $\sigma^2(t_0)$ and $\gamma(t_0)$ are given by the previous set of equations. The variance and covariance of the solutions change with a time constant that is twice as fast as the time constant with which mean changes. This is a known property of Ornstein-Uhlenbeck processes (Gardiner, 1985).

B.2.1 The time-dependent firing rate

We now investigate how the firing rate changes in response to a pulse or a step change in the input statistics. The firing rate at time t is given by $v(t) = J(\mu(t), \sigma^2(t), D_t)$ where $J(\mu, \sigma^2, D)$ is defined in (23), $\mu(t)$ and $\sigma^2(t)$ are as derived above and

$$D_t = \begin{cases} D_0 & t < 0 \\ D_1 & t \in [0, t_0] \\ D_0 & t > t_0 \end{cases}$$

is the time dependent diffusion coefficient.

We can simplify the expression to get

$$v_V(t) = \begin{cases} \dfrac{\alpha(t)}{\sqrt{\pi}} e^{-\alpha^2(t)} & t < 0 \\[2ex] \left(\dfrac{D_1}{e^{-2t} D_0 + (1 - e^{-2t}) D1} \right) \dfrac{\alpha(t)}{\sqrt{\pi}} e^{-\alpha^2(t)} & t \in [0, t_0] \\[2ex] \left(\dfrac{D_0}{(1 + e^{-2t} - e^{-2(t - t_0)}) D_0 + (e^{-2(t - t_0)} - e^{-2t}) D_1} \right) \dfrac{\alpha(t)}{\sqrt{\pi}} e^{-\alpha^2(t)} & t > t_0 \end{cases}$$

where

$$\alpha(t) = \frac{1 - m(t)}{\sqrt{2\sigma^2(t)}}.$$

Note that $\alpha(t)$ changes continuously with t. Thus any discontinuities in the expression above are from the factors multiplying the $\frac{\alpha(t)}{\sqrt{\pi}} e^{-\alpha^2(t)}$ term. In particular, the firing rate jumps discontinuously by a factor of $\frac{D_1}{D_0}$ at time 0 and by a factor of $\frac{D_1}{e^{-2t_0} D_0 + (1 - e^{-2t_0}) D_1}$ at time t_0. If we change the mean of the input signal, but do not change the variance of the input signal (by setting $\mu_0 \neq \mu_1$ and $D_0 = D_1$), then the firing rate changes continuously with time constant $\frac{1}{g} = 1$. If, instead, we change D and keep μ constant (by setting $\mu_0 = \mu_1$ and $D_0 \neq D_1$), the firing rate has jump discontinuities at time 0 and t_0, and changes with a faster time constant of $\frac{1}{2g} = \frac{1}{2}$.

B.2.2 The time dependent cross-covariance

We now look at the effects of changes in the input parameters on the conditional firing rate, $v_{V|W}(\tau, t)$. We first derive look at lag $\tau = 0$. The quantity $v_{V|W}(0, t)$ quantifies the tendency of the neurons to fire together. The conditional distribution, $P(t, V | W(t) = 1)$, of $V(t)$ given that W crossed threshold at time t is a Gaussian with mean and variance given respectively by

$$m_c(0, t) = m(t) + \rho(t)(1 - m(t))$$

and

$$\sigma_c^2(0, t) = \sigma^2(t)(1 - \rho(t))$$

where $\rho(t) = \frac{\gamma(t)}{\sigma^2(t)}$ is the sub-threshold correlation and $m(t)$, $\sigma^2(t)$, and $\gamma(t)$ are derived in the previous subsection. The firing rate at lag $\tau = 0$ is then given by

$$v_{V|W}(0, t) = v(m_c(0, t), \sigma_c^2(0, t), D_t).$$

We now derive the conditional firing rate for times $t > 0$ and lags $\tau > 0$. We break the derivation into three cases. The distribution of $V(t + \tau)$ conditioned on a spike in W at time t is a one dimensional Gaussian. If $t + \tau < t_0$, the

mean and variance of this Gaussian are given by

$$\left.\begin{array}{ll} m_c(\tau,t) & = e^{-\tau}m_c(0,t)+(1-e^{-\tau})\mu_1 \\ \sigma_c^2(\tau,t) & = e^{-2\tau}\sigma_c^2(0,t)+(1-e^{-2\tau})D_1 \end{array}\right\} t \in [0,t_0],\ t+\tau \le t_0$$

If $t \in [t,t_0]$, but $t+\tau > t_0$, the mean and variance are

$$\left.\begin{array}{ll} m_c(\tau,t) & = e^{-((t+\tau)-t_0)}m_c(t_0-t,t)+(1-e^{-((t+\tau)-t_0)})\mu_0 \\ \sigma_c^2(\tau,t) & = e^{-2((t+\tau)-t_0)}\sigma_c^2(t_0-t,t)+(1-e^{-2((t+\tau)-t_0)})D_0 \end{array}\right\} t \in [0,t_0],\ t+\tau > t_0.$$

Finally, when $t > t_0$, the mean and variance are

$$\left.\begin{array}{ll} m_c(\tau,t) & = e^{-\tau}m_c(0,t)+(1-e^{-\tau})\mu_0 \\ \sigma_c^2(\tau,t) & = e^{-2\tau}\sigma_c^2(0,t)+(1-e^{-2\tau})D_0 \end{array}\right\} t > t_0.$$

The conditional firing rate is then given by

$$v_{V|W}(\tau,t) = J(m_c(\tau,t),\sigma_c^2(\tau,t),D_{t+\tau}).$$

Acknowledgements

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